## **Research Article**



# *ATM* rs189037 significantly increases the risk of cancer in non-smokers rather than smokers: an updated meta-analysis

#### Xiaoxia He<sup>1,\*</sup>, Peng Wang<sup>2,\*</sup>, Ying Li<sup>1</sup> and () Na Shen<sup>1</sup>

<sup>1</sup>Department of Laboratory Medicine, Tongji Hospital, Tongji Medical College, Huazhong University of Science and Technology, Wuhan 430030, China; <sup>2</sup>Institute and Department of Infectious Disease, Tongji Hospital, Tongji Medical College, Huazhong University of Science and Technology, Wuhan, China

Correspondence: Na Shen (shenna@tjh.tjmu.edu.cn)



Rs189037 (G>A) is an important functional variant with ataxia telangiectasia mutated (ATM) gene, which might affect ATM's expression involvement in several human cancers. Increasing evidence reveals that smoking-related cancers have distinct molecular characteristics from non-smoking cancers. Until now, the role of ATM rs189037 in cancer risk stratified by smoking status still remains unclear. To evaluate the association between ATM rs189037 and cancer risk based on smoking status, we performed this meta-analysis by a comprehensive literature search via databases of PubMed, Embase, Web of Science and CNKI, updated till January 2019. Multivariate-adjusted odds ratios (ORs) and 95% confidence intervals (CIs) were extracted from eligible studies if available, to assess the relationship strengths. A total of seven eligible studies were included, comprising 4294 cancer patients (smokers: 1744 [40.6%]) and 4259 controls (smokers: 1418 [33.3%]). Results indicated a significant association of ATM rs189037 with cancer risk. In non-smokers, compared with GG genotype, AA genotype increased a 1.40-fold risk of overall cancer (OR = 1.40, 95% CI = 1.15-1.70, $P_{\text{heterogeneity}} = 0.433, I^2 = 0.0\%$ ). Subgroup analysis in lung cancer (LC) also exhibited a significant result (OR = 1.41, 95% CI = 1.15-1.73, P<sub>heterogeneity</sub>=0.306, I<sup>2</sup> = 17.0%) only in non-smokers. However, the association was not observed in smokers, no matter for overall cancer or for LC. Our findings highlight that ATM rs189037 significantly increases cancer susceptibility in non-smokers, rather than in smokers. The association is prominent in LC.

# Introduction

The incidence and mortality of human cancer are rapidly increasing, with an estimation of 18.1 million new cases and 9.6 million deaths in 2018 worldwide [1]. It is expected to be the leading cause of death and the single greatest threat to life expectancy in the 21st century [2]. Several environmental factors are revealed to play a role in carcinogenesis, including air pollutants, alcoholism, and virus infection [3,4]. Smoking is estimated to account for more than 30% of all cancer deaths and 90% of lung cancer (LC) deaths, and approximately 62% of all recently diagnosed cancer patients are reported as smokers, which is becoming the most prominent risk factor for human cancer [5].

Besides environmental risk factors, genetic predisposition is also crucial for occurrence and development of cancer [6–8]. Researches have uncovered many candidate genes associated with tumorigenesis, such as DNA damage checkpoint genes (DDCGs). As a famous member of DDCG, *ataxia telangiectasia mutated* (*ATM*) encodes a serine/threonie protein kinase to play a major role in cell cycle checkpoints and DNA repair initiation by phosphorylating some key factors (e.g. p53), which is frequently mutated

\*These authors contributed equally to this work.

Received: 02 May 2019 Revised: 07 June 2019 Accepted: 13 June 2019

Accepted Manuscript Online: 14 June 2019 Version of Record published: 28 June 2019



in human cancers [7,9]. ATM rs189037 (G>A), located at the 5'UTR of its promoter, is an important variant reportedly involving susceptibility to several cancers, but results remain inconclusive [10–16].

Some meta-analyses have made efforts to evaluate the role of *ATM* rs189037 in cancer risk. An early one was performed by Kang et al. (2014) [17], which only included one case–control study about rs189037. Then, Bhowmik et al. (2015) [18] and Yan et al. (2017) [19] conducted such analyses respectively focusing on specific cancer types. The latest meta-analysis was published by Zhao et al. (2019) [20], which explored the association between rs189037 and all cancer risk. However, all these meta-analyses did not consider effects of smoking, a most important environmental risk factor affecting most types of cancer. Additionally, they pooled results simply based on genotype information, rather than using confounder-adjusted odds ratio (OR), which possibly induced some bias from original studies. Therefore, we carried out this updated meta-analysis, aiming to use more refined data to clarify the effects of *ATM* rs189037 on cancer risk stratified by smoking status.

# Materials and methods

This meta-analysis was carried out according to the statement of the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) [21].

#### Literature search and eligibility criteria

Multiple databases including PubMed, Web of Science, Embase, and CNKI were searched for available relevant studies, without any restriction (updated till January 2019). The search items were used as follows: 'ATM', 'polymorphism', 'variant', 'cancer', 'smoking', 'cigarette' and 'rs189037'. We also performed manual search by reviewing the reference lists of identified publications for potentially relevant studies.

A study was considered eligible if it met all the following criteria: (i) it was a case-control study to investigate the association between *ATM* rs189037 and cancer risk; (ii) it reported the OR and 95% confidence interval (CI), or provided allele frequency and/or genotype distribution of *ATM* rs189037 in cases and controls; and (iii) it evaluated the effects of *ATM* rs189037 on cancer risk stratified by smoking status. If authors published multiple articles based on the same or overlapping datasets, we chose the study with the largest sample size. Exclusion criteria were as follows: (i) review, meta-analysis, comment, conference abstract, or experimental research; and (ii) articles without healthy controls or with duplication of earlier studies. Two independent authors conducted the literature search and study selection and discrepancy was solved by discussion.

#### Data extraction and quality evaluation

Two authors independently extracted the items from each included study, including the first author's name, publication year, cancer type, country, ethnicity, the number of cases and controls, genotyping methods, and proportion of males and smokers. In addition, multivariate-adjusted OR and 95% CI, genotype distribution, and allele frequency based on smoking status were also recorded from these eligible studies.

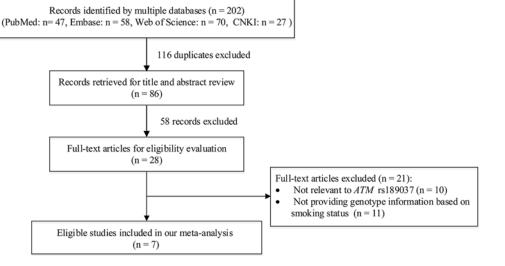
The quality of each included study was evaluated by the Newcastle–Ottawa scale (NOS), with scores in a range from '0' to '9' [22]. Quality evaluation was not an exclusion criterion for eligible studies (Supplementary Table S1).

### **Statistical analysis**

The Hardy–Weinberg equilibrium was assessed in genotypes of controls by using a  $\chi^2$  test. The strength of the association between *ATM* rs189037 and cancer risk was measured with OR and 95% CI. Multivariate-adjusted ORs and 95% CIs were preferentially extracted from included studies if available, otherwise unadjusted ORs and 95% CIs were calculated based on genotypes or allele frequencies. Cochran's Q test and  $I^2$  statistic were used to evaluate the heterogeneity among studies, and P < 0.10 or  $I^2 > 50$ % indicates significant heterogeneity [23]. A random-effects model was applied to pool results under significant heterogeneity, otherwise a fixed-effects model was used [24]. Moreover, subgroup analysis was carried out to further explore more specific roles of *ATM* rs189037 in cancer risk. We conducted one-way sensitivity analysis to assess the stability of pooled results. In addition, we also examined publication bias by Begg's and Egger's tests [25,26]. A two-sided  $P \le 0.05$  was considered as significant, unless otherwise specified. Our meta-analysis was performed by Stata 12.0 software (College Station, TX, U.S.A.).

# **Results** Characteristics of included studies

Initially, we identified 202 records from a comprehensive search via different databases. After removing 116 duplicates, we also excluded 58 records by reviewing titles and abstracts due to not being original articles (e.g. review,





#### Table 1 Characteristics of included studies

| Study              | Туре | Country | Ethnicity   | Cases/Controls | Male<br>(case/control), <i>n</i><br>(%) | Smokers<br>(case/control), <i>n</i><br>(%) | Genotyping<br>method | NOS<br>score |
|--------------------|------|---------|-------------|----------------|---|--|----------------------|--------------|
| Bau et al. (2010)  | OC   | China   | East Asians | 620/620        | 586 (94.5)/582 (93.9)                   | 458 (73.9)/443 (71.5)                      | PCR-RFLP             | 5            |
| Lo et al. (2010)   | LC   | China   | East Asians | 730/730        | 384 (52.6)/384 (52.6)                   | 268 (36.7)/268 (36.7)                      | MassARRAY            | 5            |
| Liu et al. (2014)  | LC   | China   | East Asians | 852/852        | 485 (56.9)/490 (57.5)                   | 477 (66.0)/273 (32.0)                      | TaqMan assay         | 6            |
| Shen et al. (2014) | LC   | China   | East Asians | 487/516        | All females                             | All non-smokers                            | TaqMan assay         | 7            |
| Yu et al. (2015)   | ESCC | China   | East Asians | 303/304        | 258 (85.1)/253 (83.2)                   | 214 (70.60)/153 (50.3)                     | TaqMan assay         | 6            |
| Han et al. (2017)  | LC   | China   | East Asians | 181/181        | 61 (33.7)/61 (33.7)                     | All non-smokers                            | MassARRAY            | 5            |
| Wang et al. (2018) | CRC  | China   | East Asians | 1121/1056      | 631 (56.3)/561 (53.1)                   | 327 (29.2)/281 (26.6)                      | TaqMan assay         | 6            |

Abbreviation: PCR-RFLP, polymerase chain reaction and restriction fragment length polymorphism.

meta-analysis, comment), and not related to cancer risk. Out of the remaining 28 records for full-text review, we further removed 21 studies based on the following reasons: (i) relevant to other variants of *ATM* but not rs189037; and (ii) providing insufficient genotype information of *ATM* rs189037 based on smoking status. Finally, a total of seven eligible studies were included for meta-analysis (Figure 1) [10-16].

Characteristics of these studies are summarized in Table 1. All the subjects were East Asians from China. Overall, there were 2405 males (56.0%) and 1744 smokers (40.6%) in 4294 cancer patients, and were 2331 males (54.7%) and 1418 smokers (33.3%) in 4259 controls. Four studies focused on LC [11–13,15], and the remaining studies focused on oral cancer (OC) [10], esophageal squamous cell carcinoma (ESCC) [14] and colorectal cancer (CRC) [16], respectively. All these studies had an NOS score  $\geq$  5. Table 2 shows the genotype distribution and allele frequency of *ATM* rs189037 in smokers, non-smokers, and overall subjects.

# Association between *ATM* rs189037 and cancer risk stratified by smoking status

The allelic, dominant, recessive, and codominant models were applied to pool results (Table 3). Overall, *ATM* rs189037 A allele exhibited a 1.17-fold increased risk of cancer compared with the G allele (OR = 1.17, 95% CI = 1.0–1.30). Other genetic models showed the same results. Analysis based on smoking status, we further found a consistent and significant association of *ATM* rs189037 with cancer risk in non-smokers (allelic model: OR = 1.16, 95% CI = 1.05–1.28; dominant model: OR = 1.43, 95% CI = 1.22–1.69; recessive model: OR = 1.14, 95% CI = 1.01–1.29; codominant AA vs GG model: OR = 1.40, 95% CI = 1.15–1.70; additive model: OR = 1.16, 95% CI = 1.06–1.28). Interesting, however, the association was not observed in smokers (all *P*>0.05).

| Study              | Smoking<br>exposure   | Genotype (GG/GA/AA)  |                      | Minor allele frequency (A allele) |              | P <sub>HWE</sub> |
|--------------------|-----------------------|----------------------|----------------------|-----------------------------------|--------------|------------------|
|                    |                       | Cases                | Controls             | Cases (%)                         | Controls (%) |                  |
| Bau et al. (2010)  | Overall               | 181/277/162          | 239/285/96           | 48.47                             | 38.47        | 0.470            |
|                    | Smokers               | 337/121 <sup>1</sup> | 374/69 <sup>1</sup>  | -                                 | -            | -                |
|                    | Non-smokers           | 121/41 <sup>1</sup>  | 150/27 <sup>1</sup>  | -                                 | -            | -                |
| Lo et al. (2010)   | Overall               | 238/345/145          | 239/354/124          | 43.61                             | 41.98        | 0.717            |
|                    | Smokers               | 103/122/42           | 82/131/49            | 38.58                             | 43.70        | 0.794            |
|                    | Non-smokers           | 135/223/103          | 157/223/72           | 46.53                             | 40.60        | 0.626            |
| Liu et al. (2014)  | Overall               | 217/435/200          | 264/434/154          | 49.00                             | 43.54        | 0.293            |
|                    | Smokers               | 120/249/108          | 87/129/57            | 48.74                             | 44.51        | 0.473            |
|                    | Non-smokers           | 97/186/92            | 177/305/97           | 49.33                             | 43.09        | 0.075            |
| Shen et al. (2014) | Overall (non-smokers) | 148/240/99           | 152/272/92           | 44.97                             | 44.19        | 0.119            |
| Yu et al. (2015)   | Overall               | 106/139/58           | 114/145/45           | 42.08                             | 38.65        | 0.920            |
|                    | Smokers               | 72/97/45             | 59/67/27             | 43.69                             | 39.54        | 0.298            |
|                    | Non-smokers           | 34/42/13             | 55/78/18             | 38.20                             | 37.75        | 0.223            |
| Han et al. (2017)  | Overall (non-smokers) | 56/83/39             | 54/92/32             | 45.22                             | 43.82        | 0.507            |
| Wang et al. (2018) | Overall               | 336/543/227          | 362/491/191          | 45.07                             | 41.81        | 0.280            |
|                    | Smokers               | 107/213 <sup>2</sup> | 106/171 <sup>2</sup> | -                                 | -            | -                |
|                    | Non-smokers           | 229/557 <sup>2</sup> | 256/511 <sup>2</sup> | -                                 | -            | -                |

#### Table 2 Genotype distribution and allele frequency of ATM rs189037 stratified by smoking status

<sup>2</sup>indicates the number of GG/(GA+AA).

#### Table 3 Meta-analysis for the association between ATM rs189037 and cancer risk stratified by smoking status

| Genetic model <sup>1</sup>      | Effect size |                  | Heterogeneity  |                    | Publication bias  |                    |
|---------------------------------|-------------|------------------|----------------|--------------------|-------------------|--------------------|
|                                 | n           | OR (95% CI)      | Pheterogeneity | l <sup>2</sup> (%) | P <sub>Begg</sub> | P <sub>Egger</sub> |
| Allelic model                   |             |                  |                |                    |                   |                    |
| Overall                         | 7           | 1.17 (1.06–1.30) | 0.026          | 58.0               | 0.764             | 0.738              |
| Non-smokers                     | 5           | 1.16 (1.05–1.28) | 0.324          | 14.2               | 0.806             | 0.514              |
| Smokers                         | 3           | 1.04 (0.81–1.34) | 0.044          | 68.0               | 0.956             | 0.602              |
| Dominant model                  |             |                  |                |                    |                   |                    |
| Overall                         | 7           | 1.32 (1.19–1.47) | 0.130          | 39.2               | 0.548             | 0.780              |
| Non-smokers                     | 6           | 1.43 (1.22–1.69) | 0.642          | 0.0                | 0.707             | 0.894              |
| Smokers                         | 4           | 1.24 (0.84–1.82) | 0.013          | 72.1               | 0.734             | 0.403              |
| Recessive model                 |             |                  |                |                    |                   |                    |
| Overall                         | 7           | 1.19 (1.08–1.30) | 0.107          | 42.5               | 0.548             | 0.344              |
| Non-smokers                     | 6           | 1.14 (1.01–1.29) | 0.504          | 0.0                | 0.851             | 0.263              |
| Smokers                         | 4           | 1.12 (0.83–1.50) | 0.048          | 62.0               | 0.308             | 0.902              |
| Codominant model<br>(AA vs. GG) |             |                  |                |                    |                   |                    |
| Overall                         | 7           | 1.42 (1.18–1.70) | 0.077          | 47.4               | 0.881             | 0.727              |
| Non-smokers                     | 5           | 1.40 (1.15–1.70) | 0.433          | 0.0                | 1.000             | 0.608              |
| Smokers                         | 3           | 1.12 (0.68–1.83) | 0.064          | 63.6               | 0.602             | 0.983              |
| Codominant model<br>(GA vs. GG) |             |                  |                |                    |                   |                    |
| Overall                         | 7           | 1.11 (1.00-1.22) | 0.338          | 11.9               | 0.548             | 0.153              |
| Non-smokers                     | 5           | 1.01 (0.86–1.18) | 0.713          | 0.0                | 1.000             | 0.391              |
| Smokers                         | 3           | 1.06 (0.72–1.57) | 0.063          | 63.8               | 0.602             | 0.852              |
| Additive model                  |             |                  |                |                    |                   |                    |
| Overall                         | 7           | 1.17 (1.07–1.29) | 0.044          | 53.6               | 0.764             | 0.655              |
| Non-smokers                     | 5           | 1.16 (1.06–1.28) | 0.317          | 15.3               | 0.806             | 0.515              |
| Smokers                         | 3           | 1.04 (0.81–1.34) | 0.046          | 67.5               | 0.602             | 0.937              |

<sup>1</sup>Allelic model refers to A allele vs. G allele; dominant model refers to AA+GA vs. GG; recessive model refers to AA vs. GG+GA.



Considering that smoking is the crucial pathogenic factor for LC, we further evaluated effects of *ATM* rs189037 on LC based on smoking status (Figure 2). Overall, *ATM* rs189037 AA carriers had more risk of LC than wild-type carriers (OR = 1.33, 95% CI = 1.11–1.58). Specially, the association was more notable in non-smokers (OR = 1.41, 95% CI = 1.15–1.73). Also, we did not find any association of this variant with LC risk in smokers. In addition, individuals with GA genotype did not suffer more susceptibility to overall cancer or LC than those with GG genotype, no matter in smokers or in non-smokers (all P > 0.05).

#### Sensitivity analysis and publication bias

To assess the stability of pooled results, we conducted one-way sensitivity analysis by excluding one study at a time. It demonstrated that our pooed results were quite stable in both non-smokers and smokers (Supplementary Figure S1). In addition, results from Begg's and Egger's tests showed that there was no obvious publication bias in our meta-analyses under all genetic models (Table 3).

# Discussion

In this meta-analysis, we have four findings as follows: (i) *ATM* rs189037 significantly increased the overall risk of cancer under most of genetic models; (ii) the risky role of *ATM* rs189037 was prominent in non-smokers, but not observed in smokers; (iii) results focusing on LC were consistent with results of overall cancer; and (iv) GA genotype carriers of *ATM* rs189037 appeared not to suffer more cancer risk than GG wild-type carriers, no matter in smokers or in non-smokers.

In agreement with previous meta-analyses [17–20], our work suggested that *ATM* rs189037 is a risky variant for cancer susceptibility, but we have some new highlights. First, we elucidated the association between this variant and cancer risk stratified by smoking status, and specially focused on LC. Results demonstrated different effects of *ATM* rs189037 in non-smokers and in smokers. Second, we pooled results by extracting multivariate-adjusted ORs and 95% CIs if available, greatly reducing confounder bias from original studies. Third, we performed a comprehensive literature and included a study ignored by previous meta-analyses [14].

Rs189037 (G>A) is located at the promoter region of ATM gene, alleles of which may possibly have different binding affinities to transcription factors (e.g. AP-2 $\alpha$ ) or change ATM folding structure to affect its mRNA expression, and AA genotype was reported to show a lower ATM expression than GG genotype [16,27,28]. Reduced ATM expression may impair its normal function, lead to uncontrolled cell cycle, abnormal DNA repair and apoptosis, and finally increase the susceptibility to cancer. This could well explain our findings that individuals with ATM rs189037 suffer more cancer risk. Furthermore, our results demonstrated the association was significant in non-smokers rather than in smokers, no matter for overall cancer or for LC. As the most important risk factor for human cancer, smoking induces a serious of potential carcinogens to generate DNA damage and oxidative stress, resulting in gene mutations and genomic instability. Smoking-related cancers have a high mutational load and highly significant molecular heterogeneity [29]. Compared with non-smokers, smokers probably alert more gene pathways to remove those tobacco-induced DNA adducts and activate more antioxidant mechanisms to fight against smoking-related stress. Consistently, we identified obvious heterogeneity among included studies in smokers, whereas we observed a good homogeneity among studies in non-smokers, both for overall cancer and LC (all  $P_{\text{heterogeneity}} > 0.10$ ,  $I^2 < 50\%$ ). These results suggest that smoking-related cancers has distinct molecular characteristics from non-smoking-related cancers. Supportively, a recent study revealed that LCr in smokers and in non-smokers showed quite different tumor immune microenvironments [30].

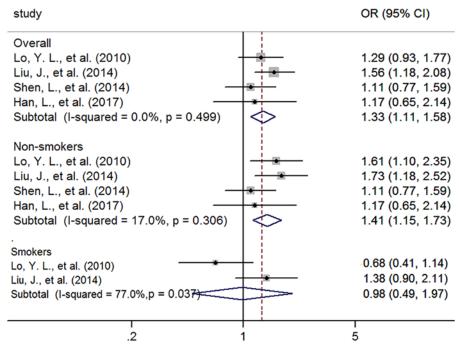
However, some limitations should be acknowledged here. First, only seven studies were eligible for inclusion in our study. The small number of studies possibly affected the conclusion's extrapolation to some extent. More studies are still needed to verify our results. Second, we did not evaluate the effects of this variant in other ethnic populations, since all included subjects were Chinese. At last, except for LC, number of studies focusing on other cancer is limited, so we could not explore its role based on subgroup analysis in other cancer types.

# Conclusions

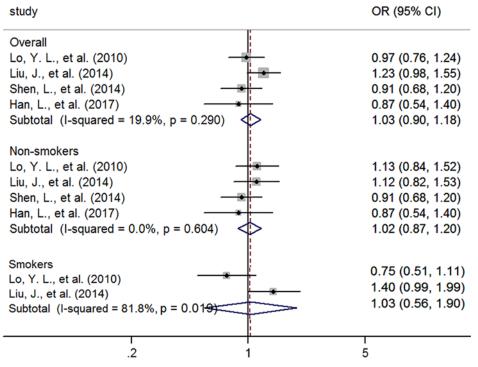
Our study highlights that there is a significant association between *ATM* rs189037 and cancer risk in non-smokers, rather than in smokers. This association is prominent in LC. Our work not only provides a new insight into the pathogenic role of *ATM* variants in occurrence of cancer, but also supports the distinct molecular characteristics of cancers between smokers and non-smokers. More studies are still needed to verify our results in the future.



# (A) Lung cancer: AA vs GG



# (B) Lung cancer: GA vs GG



#### Figure 2. Forest plots of the association between ATM rs189037 and LC risk

Forest plots for evaluation of the association between ATM rs189037 and LC risk under the codominant models of AA vs GG (A) and GA vs GG (B).



#### Author Contribution

N.S. designed the study and revised the manuscript. X.H. and P.W. performed literature search, collected data, performed meta-analysis and wrote the manuscript. Y.L. performed literature search and collected data. All authors read and approved the final manuscript.

#### **Competing Interests**

The authors declare that there are no competing interests associated with the manuscript.

#### Funding

The authors declare that there are no sources of funding to be acknowledged.

#### Availability of data and material

All data in the present study are included in this published article and the additional files.

#### Abbreviations

ATM, *ataxia telangiectasia mutated*; CI, confidence interval; DDCG, DNA damage checkpoint gene; LC, lung cancer; NOS, Newcastle–Ottawa scale; OR, odds ratio.

#### References

- 1 Ferlay, J., Colombet, M., Soerjomataram, I., Mathers, C., Parkin, D.M., Pineros, M. et al. (2019) Estimating the global cancer incidence and mortality in 2018: GLOBOCAN sources and methods. *Int. J. Cancer* **144**, 1941–1953
- 2 Bray, F., Ferlay, J., Soerjomataram, I., Siegel, R.L., Torre, L.A. and Jemal, A. (2018) Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. *CA Cancer J. Clin.* **68**, 394–424, https://doi.org/10.3322/caac.21492
- 3 Parkin, D.M., Boyd, L. and Walker, L.C. (2011) 16. The fraction of cancer attributable to lifestyle and environmental factors in the UK in 2010. *Br. J. Cancer* **105**, S77–S81, https://doi.org/10.1038/bjc.2011.489
- 4 Islami, F., Chen, W., Yu, X.Q., Lortet-Tieulent, J., Zheng, R., Flanders, W.D. et al. (2017) Cancer deaths and cases attributable to lifestyle factors and infections in China, 2013. Ann. Oncol. 28, 2567–2574
- 5 Karam-Hage, M., Cinciripini, P.M. and Gritz, E.R. (2014) Tobacco use and cessation for cancer survivors: an overview for clinicians. *CA Cancer J. Clin.* **64**, 272–290, https://doi.org/10.3322/caac.21231
- 6 Shen, N., Lu, Y., Wang, X., Peng, J., Zhu, Y. and Cheng, L. (2017) Association between rs2853669 in TERT gene and the risk and prognosis of human cancer: a systematic review and meta-analysis. *Oncotarget* **8**, 50864–50872
- 7 Negrini, S., Gorgoulis, V.G. and Halazonetis, T.D. (2010) Genomic instability–an evolving hallmark of cancer. *Nat. Rev. Mol. Cell Biol.* **11**, 220–228, https://doi.org/10.1038/nrm2858
- 8 Y-j, Lu, Wang, P., Peng, J., Wang, X., Zhu, Y.-w. and Shen, N. (2017) Meta-analysis reveals the prognostic value of circulating tumour cells detected in the peripheral blood in patients with non-metastatic colorectal cancer. *Sci. Rep.* **7**, 905, https://doi.org/10.1038/s41598-017-01066-y
- 9 Choi, M., Kipps, T. and Kurzrock, R. (2016) ATM mutations in cancer: therapeutic implications. *Mol. Cancer Ther.* **15**, 1781–1791, https://doi.org/10.1158/1535-7163.MCT-15-0945
- 10 Bau, D.T., Chang, C.H., Tsai, M.H., Chiu, C.F., Tsou, Y.A., Wang, R.F. et al. (2010) Association between DNA repair gene ATM polymorphisms and oral cancer susceptibility. *Laryngoscope* **120**, 2417–2422, https://doi.org/10.1002/lary.21009
- 11 Lo, Y.L., Hsiao, C.F., Jou, Y.S., Chang, G.C., Tsai, Y.H., Su, W.C. et al. (2010) ATM polymorphisms and risk of lung cancer among never smokers. *Lung Cancer* 69, 148–154, https://doi.org/10.1016/j.lungcan.2009.11.007
- 12 Liu, J., Wang, X., Ren, Y., Li, X., Zhang, X. and Zhou, B. (2014) Effect of single nucleotide polymorphism Rs189037 in ATM gene on risk of lung cancer in Chinese: a case-control study. *PLoS ONE* 9, e115845, https://doi.org/10.1371/journal.pone.0115845
- 13 Shen, L., Yin, Z., Wu, W., Ren, Y., Li, X. and Zhou, B. (2014) Single nucleotide polymorphism in ATM gene, cooking oil fumes and lung adenocarcinoma susceptibility in Chinese female non-smokers: a case-control study. *PLoS ONE* **9**, e96911, https://doi.org/10.1371/journal.pone.0096911
- 14 Yu, Q., Qiu, H. and Xiong, H. (2015) HPV seropositivity synergized with ATM variants increases the risk of esophageal squamous cell carcinoma in drinkers. *Acta Med. Univ. Sci. Technol.* 44, 052–059
- 15 Han, L., Lee, C.K., Pang, H., Chan, H.T., Lo, I.L., Lam, S.K. et al. (2017) Genetic predisposition to lung adenocarcinoma among never-smoking Chinese with different epidermal growth factor receptor mutation status. *Lung Cancer* **114**, 79–89, https://doi.org/10.1016/j.lungcan.2017.10.012
- 16 Wang, S., Zhang, Y., Chen, M., Wang, Y., Feng, Y., Xu, Z. et al. (2018) Association of genetic variants in ATR-CHEK1 and ATM-CHEK2 pathway genes with risk of colorectal cancer in a Chinese population. *Oncotarget* **9**, 26616–26624
- 17 Kang, J., Deng, X.Z., Fan, Y.B. and Wu, B. (2014) Relationships of FOXE1 and ATM genetic polymorphisms with papillary thyroid carcinoma risk: a meta-analysis. *Tumour Biol.* **35**, 7085–7096, https://doi.org/10.1007/s13277-014-1865-5
- 18 Bhowmik, A., Nath, S., Das, S., Ghosh, S.K. and Choudhury, Y. (2015) ATM rs189037 (G > A) polymorphism and risk of lung cancer and head and neck cancer: a meta-analysis. *Meta Gene* 6, 42–48
- 19 Yan, Z., Tong, X., Ma, Y., Liu, S., Yang, L., Yang, X. et al. (2017) Association between ATM gene polymorphisms, lung cancer susceptibility and radiation-induced pneumonitis: a meta-analysis. *BMC Pulmon. Med.* **17**, 205, https://doi.org/10.1186/s12890-017-0555-7





- 20 Zhao, Z.L., Xia, L., Zhao, C. and Yao, J. (2019) ATM rs189037 (G > A) polymorphism increased the risk of cancer: an updated meta-analysis. *BMC Med. Genet.* 20, 28
- 21 Moher, D., Liberati, A., Tetzlaff, J. and Altman, D.G. (2009) Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. *PLoS Med.* **6**, e1000097, https://doi.org/10.1371/journal.pmed.1000097
- 22 Stang, A. (2010) Critical evaluation of the Newcastle-Ottawa scale for the assessment of the quality of nonrandomized studies in meta-analyses. *Eur. J. Epidemiol.* **25**, 603–605, https://doi.org/10.1007/s10654-010-9491-z
- 23 Higgins, J.P. and Thompson, S.G. (2002) Quantifying heterogeneity in a meta-analysis. Stat. Med. 21, 1539–1558, https://doi.org/10.1002/sim.1186
- 24 Harris, R.J., Bradburn, M.J., Deeks, J.J., Harbord, R.M., Altman, D.G. and Sterne, J.A.C. (2008) metan: fixed- and random-effects meta-analysis. *Stata J.* **8**, 3–28, https://doi.org/10.1177/1536867X0800800102
- 25 Begg, C.B. and Mazumdar, M. (1994) Operating characteristics of a rank correlation test for publication bias. *Biometrics* **50**, 1088–1101, https://doi.org/10.2307/2533446
- 26 Egger, M., Davey Smith, G., Schneider, M. and Minder, C. (1997) Bias in meta-analysis detected by a simple, graphical test. *BMJ* **315**, 629–634, https://doi.org/10.1136/bmj.315.7109.629
- 27 Chen, T., Dong, B., Lu, Z., Tian, B., Zhang, J., Zhou, J. et al. (2010) A functional single nucleotide polymorphism in promoter of ATM is associated with longevity. *Mech. Ageing Dev.* **131**, 636–640, https://doi.org/10.1016/j.mad.2010.08.009
- 28 Zhang, L., Yang, M., Bi, N., Fang, M., Sun, T., Ji, W. et al. (2010) ATM polymorphisms are associated with risk of radiation-induced pneumonitis. *Int. J. Radiat. Oncol. Biol. Phys.* 77, 1360–1368, https://doi.org/10.1016/j.ijpobp.2009.07.1675
- 29 Kytola, V., Topaloglu, U., Miller, L.D., Bitting, R.L., Goodman, M.M., D'Agostino, R.B. et al. (2017) Mutational landscapes of smoking-related cancers in Caucasians and African Americans: precision oncology perspectives at Wake Forest Baptist Comprehensive Cancer Center. *Theranostics* 7, 2914–2923, https://doi.org/10.7150/thno.20355
- 30 Li, X., Li, J., Wu, P., Zhou, L., Lu, B., Ying, K. et al. (2018) Smoker and non-smoker lung adenocarcinoma is characterized by distinct tumor immune microenvironments. *Oncoimmunology* 7, e1494677, https://doi.org/10.1080/2162402X.2018.1494677